

Inference

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Inference

Definition 1: Inference

Inference is the process of drawing conclusions about a population's characteristics based on the sample data.

- We conduct **inference** by performing hypothesis tests and constructing confidence intervals.
- To make **inferences** in finite samples, we have to make assumptions regarding the sampling distribution of our estimator.

Normality of Errors

MLR Assumption 6: Normality of Errors

For each $i = 1, \dots, n$,

1. The population error u_i is independent of \mathbf{x}_i
 - ▶ This implies $\mathbb{E}[u_i | \mathbf{x}_i] = \mathbb{E}[u_i]$
2. The population error $u_i \sim \mathcal{N}(0, \sigma^2)$
 - ▶ $\mathbb{E}[u_i] = 0$ and $\mathbb{V}[u_i] = \sigma^2$
 - ▶ Could also write as $\mathbf{u} \sim \mathcal{N}(0, \sigma^2 I_n)$

- This assumption implies MLR Assumptions 4 and 5.

Conditional Normality of Outcome

Theorem 1: Conditional Normality of Outcome

Under MLR Assumptions 1-6, for each $i = 1, \dots, n$,

$$y_i | \mathbf{x}_i \sim \mathbb{N}(\mathbf{x}'_i \boldsymbol{\beta}, \sigma^2).$$

$$\mathbf{y} | X \sim \mathbb{N}(X\boldsymbol{\beta}, \sigma^2 I_n).$$

- y_i is a linear function of a normal random variable, u_i , so y_i is also normal.
- $\mathbb{E}[y_i | \mathbf{x}_i] = \mathbb{E}[\mathbf{x}'_i \boldsymbol{\beta} + u_i | \mathbf{x}_i] = \mathbf{x}'_i \boldsymbol{\beta}$.
- $\mathbb{V}[y_i | \mathbf{x}_i] = \mathbb{V}[\mathbf{x}'_i \boldsymbol{\beta} + u_i | \mathbf{x}_i] = \sigma^2$ because we treat \mathbf{x}_i and $\boldsymbol{\beta}$ as known and constant, respectively.

Normality of Errors

Question 1

Is the **normality assumption of the error term** reasonable?

Normality of Errors

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Answer to Question 1

Probably not. Some cases where it likely fails is when

1. $y_i > 0$ for all observations.
2. y_i takes on discrete values (e.g., binary outcome) for all observations.

Conditional Normality of OLS Estimator

Theorem 2: Conditional Normality of OLS Estimator

Under MLR Assumptions 1-6, for each parameter $j = 1, \dots, k$

$$\hat{\beta}_j | X \sim \mathbb{N} \left(\beta_j, \mathbb{V} \left[\hat{\beta}_j \right] \right) = \mathbb{N} \left(\beta_j, [\sigma^2 (X'X)^{-1}]_{jj} \right)$$

$$\hat{\beta} | X \sim \mathbb{N} \left(\beta, \sigma^2 (X'X)^{-1} \right).$$

Conditional Normality of OLS Estimator

Proof 1: Conditional Normality of OLS Estimator Part 1

Recall we can write $\hat{\beta}$ as

$$\hat{\beta} = \beta + (X'X)^{-1}X'u.$$

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$$\hat{\beta} = \beta + (X'X)^{-1}X'u.$$

Under MLR Assumption 6, $u \sim \mathbb{N}(0, \sigma^2 I_n)$. Since $\hat{\beta}$ is a linear combination of u , the former is also normal.

MLR Assumptions 1-4 then imply OLS is unbiased so $\mathbb{E}[\hat{\beta} | X] = \beta$ and MLR Assumptions 1-5 imply $\mathbb{V}[\hat{\beta} | X] = \sigma^2(X'X)^{-1}$. \square

Hooray!

Conditional Standard Normality of OLS Estimator

Property 1: Conditional Standard Normality of OLS Estimator

Under MLR Assumptions 1-6, since

$$\hat{\beta}_j | X \sim \mathbb{N} \left(\beta_j, [\sigma^2 (X'X)^{-1}]_{jj} \right),$$

it follows that

$$\frac{\hat{\beta}_j - \beta_j}{\sigma \sqrt{[(X'X)^{-1}]_{jj}}} \Big| X = \frac{\hat{\beta}_j - \beta_j}{\text{sd}[\hat{\beta}_j]} \Big| X \sim \mathbb{N}(0, 1).$$

- Problem: We don't know σ and, consequently, don't know the standard deviation of $\hat{\beta}_j$.

Student's t-Distribution

Theorem 3: Student's t-Distribution

The T-statistic

$$T = \frac{\hat{\beta}_j - \beta_j}{\hat{\sigma} \sqrt{[(X'X)^{-1}]_{jj}}} = \frac{\hat{\beta}_j - \beta_j}{\text{se}[\hat{\beta}_j]} \sim t_{n-k}$$

follows the **Student's t-Distribution** with $n - k$ degrees of freedom where k is the number of estimated parameters (intercept included).

- Replace σ with $\hat{\sigma}$ and we go from standard normal distribution to **Student's t-Distribution**.
- Symmetric around zero and converges to standard normal as degrees of freedom tends to infinity.

Student's t-Distribution

Question 2

What is the point of all this? Why do we care about the T -statistic? We don't know β_j so we still can't compute it. Aren't we in the same problem as with the standard normal distribution?

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What is the point of all this? Why do we care about the T -statistic? We don't know β_j so we still can't compute it. Aren't we in the same problem as with the standard normal distribution?

Answer to Question 2

Although we don't know β_j , we may be interested in the *hypothesis* that it is some value a . Then, we can plug in a for β_j in the T -statistic and carry out this hypothesis test and compute confidence intervals!

Two-Tailed T-Test

Definition 2: Two-Tailed T-Test

A **two-tailed T-Test** is a way of testing the null hypothesis H_0 that our true parameter β_j is *equal* some value a . To conduct a **two-tailed T-Test**, we:

1. Formulate a hypothesis test of $H_0 : \beta_j = a$ versus $H_A : \beta_j \neq a$.

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2. Select a significance level α .

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3. Construct the T-statistic
$$T = \frac{\hat{\beta}_j - a}{\text{se}[\hat{\beta}_j]}.$$

4. Find the t-critical value $t_{1-\frac{\alpha}{2}, n-k}$ that corresponds to $n - k$ degrees of freedom at the chosen significance level α .

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3. Construct the T-statistic $T = \frac{\hat{\beta}_j - a}{\text{se}[\hat{\beta}_j]}$.
4. Find the t-critical value $t_{1-\frac{\alpha}{2}, n-k}$ that corresponds to $n - k$ degrees of freedom at the chosen significance level α .
5. Reject the null hypothesis if and only if $|T| > t_{1-\frac{\alpha}{2}, n-k}$.

Right-Tailed T-Test

Definition 3: Right-Tailed T-Test

A **right-tailed T-Test** is a way of testing the null hypothesis H_0 that our true parameter β_j is *less than or equal* to some value a . To conduct a **right-tailed T-Test**, we:

1. Formulate a hypothesis test of $H_0 : \beta_j \leq a$ versus $H_A : \beta_j > a$.

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2. Select a significance level α .

3. Construct the T-statistic
$$T = \frac{\hat{\beta}_j - a}{\text{se}[\hat{\beta}_j]}.$$

4. Find the t-critical value $t_{1-\alpha, n-k}$ that corresponds to $n - k$ degrees of freedom at the chosen significance level α .

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2. Select a significance level α .

3. Construct the T-statistic
$$T = \frac{\hat{\beta}_j - a}{\text{se}[\hat{\beta}_j]}.$$

4. Find the t-critical value $t_{1-\alpha, n-k}$ that corresponds to $n - k$ degrees of freedom at the chosen significance level α .
5. Reject the null hypothesis if and only if $T > t_{1-\alpha, n-k}$.

Left-Tailed T-Test

Definition 4: Left-Tailed T-Test

A **left-tailed T-Test** is a way of testing the null hypothesis H_0 that our true parameter β_j is *greater than or equal* to some value a . To conduct a **left-tailed T-Test**, we:

1. Formulate a hypothesis test of $H_0 : \beta_j \geq a$ versus $H_A : \beta_j < a$.
2. Select a significance level α .

3. Construct the T-statistic
$$T = \frac{\hat{\beta}_j - a}{\text{se}[\hat{\beta}_j]}.$$

4. Find the t-critical value $t_{\alpha, n-k}$ that corresponds to $n - k$ degrees of freedom at the chosen significance level α .
5. Reject the null hypothesis if and only if $T < t_{\alpha, n-k}$.

Hypothesis Testing Terminology

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- **Confidence Level:** How sure we are of something.

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- **Null Hypothesis H_0 :** What we assume to be true.

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- **Type 1 Error (False Positive):** Rejecting a null hypothesis that is true. α corresponds to the probability we are willing to accept this error.

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- **Confidence Level:** How sure we are of something.
- **Significance Level α :** $1 - \text{Confidence Level}$.
- **Null Hypothesis H_0 :** What we assume to be true.
- **Alternative Hypothesis H_A :** What we need evidence of.
- **Type 1 Error (False Positive):** Rejecting a null hypothesis that is true. α corresponds to the probability we are willing to accept this error.
- **Type 2 Error (False Negative):** Failing to reject a null hypothesis that is false. Corresponds to the power of a test (probability of correctly rejecting a null hypothesis).

t-Critical Value

Definition 6: t-Critical Value

The **t-critical value** $t_{1-\frac{\alpha}{2}, n-k}$ for a two-tailed test is the value such that the cumulative distribution function, F_T , of a t-distributed random variable, T_{n-k} , with $n-k$ degrees of freedom equals $1 - \frac{\alpha}{2}$:

$$\begin{aligned} F_{T_{n-k}} \left(t_{1-\frac{\alpha}{2}, n-k} \right) &= \mathbb{P} \left(T_{n-k} \leq t_{1-\frac{\alpha}{2}, n-k} \right) \\ &= 1 - \frac{\alpha}{2}. \end{aligned}$$

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$$\begin{aligned} F_{T_{n-k}}(t_{1-\frac{\alpha}{2}, n-k}) &= \mathbb{P}(T_{n-k} \leq t_{1-\frac{\alpha}{2}, n-k}) \\ &= 1 - \frac{\alpha}{2}. \end{aligned}$$

- An analogous process applies for right-tailed and left-tailed tests:
 - ▶ For a right-tailed test, the t-critical value $t_{1-\alpha, n-k}$ satisfies $F_{T_{n-k}}(t_{1-\alpha, n-k}) = 1 - \alpha$.
 - ▶ For a left-tailed test, the t-critical value $t_{\alpha, n-k}$ satisfies $F_{T_{n-k}}(t_{\alpha, n-k}) = \alpha$.
 - ▶ We use the inverse CDF of the T-distribution to get $t_{1-\frac{\alpha}{2}, n-k}$.

p-value

Definition 7: p-Value

The **p-value** p corresponding to a test statistic is the smallest level of significance at which the null hypothesis H_0 can be rejected.

p-value for T-Test

Example 1: p-value for Two-Tailed T-Test

The p-value p corresponding to a two-tailed test T-test is given by

$$\begin{aligned} p &= 2 * \mathbb{P}(T_{n-k} > |T|) \\ &= 2 * (1 - F_{T_{n-k}}(|T|)). \end{aligned}$$

where T_{n-k} is a t -distributed random variable with degrees of freedom $n - k$.

- We multiply by two because we are doing a two-tailed test and need the area of both the upper and lower tails.
 - ▶ The area in these lower and upper tails is identical because the t -distribution is symmetric around zero.

Economic Versus Statistical Significance

Definition 8: Economic Significance

We say a parameter estimate $\hat{\beta}_j$ is **economically significant** if it is sufficiently large.

Economic Versus Statistical Significance

Question 3

Suppose we estimate the model

$$\ln(\text{income}_i) = \beta_0 + \beta_1 \text{education}_i + \beta_2 \text{experience}_i + u_i$$

where *income* is a person's yearly income, *education* is measured in years, and *experience* is measured in years.

Suppose $\hat{\beta}_1 = 0.01$ with a p-value of 0.002 and $\hat{\beta}_2 = 0.3$ with a p-value of 0.045. Are these estimates **statistically significant** at the $\alpha = 0.05$ level. What about **economically significant**?

Economic Versus Statistical Significance

Answer to Question 3

1. $\hat{\beta}_1$ is statistically significant at the $\alpha = 0.05$ level because $0.002 < 0.05$. However, it is not economically significant because our model predicts that the effect of an additional year of education only amounts to a person's yearly income increasing by 1%.
2. $\hat{\beta}_2$ is statistically significant at the $\alpha = 0.05$ level because $0.045 < 0.05$. It is economically significant because our model predicts an additional year of education amounts to a person's yearly income increasing by 30% which is quite a big amount.

Confidence Intervals

Theorem 4: Confidence Interval Estimator

The $100(1 - \alpha)\%$ confidence interval for a parameter β_j is given by

$$\left[\hat{\beta}_j - t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}(\hat{\beta}_j), \hat{\beta}_j + t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}(\hat{\beta}_j) \right].$$

- Same as saying $\hat{\beta}_j \pm t_{1-\frac{\alpha}{2}, n-k} * \text{se}[\hat{\beta}_j]$.
- Gives us a region of values we can say the true β_j lies in with $100(1 - \alpha)\%$ confidence.
- Confidence interval widens as α increases.

Confidence Intervals

Proof 2: Confidence Interval Estimator Part 1

Let $t_{1-\frac{\alpha}{2}, n-k}$ be the critical value for the t-distribution with $n - k$ degrees of freedom. Then, the probability that a t-distributed random variable T_{n-k} with $n - k$ degrees of freedom is greater than $t_{1-\frac{\alpha}{2}, n-k}$ is $1 - F_{T_{n-k}}(t_{1-\frac{\alpha}{2}, n-k}) = 1 - (1 - \frac{\alpha}{2}) = \frac{\alpha}{2}$. By the symmetry of the t-distribution, we have $F_{T_{n-k}}(-t_{1-\frac{\alpha}{2}, n-k}) = F_{T_{n-k}}(t_{\frac{\alpha}{2}, n-k}) = \frac{\alpha}{2}$. Consequently,

$$\begin{aligned} & \mathbb{P}(-t_{1-\frac{\alpha}{2}, n-k} \leq T_{n-k} \leq t_{1-\frac{\alpha}{2}, n-k}) \\ &= F_{T_{n-k}}(t_{1-\frac{\alpha}{2}, n-k}) - F_{T_{n-k}}(-t_{1-\frac{\alpha}{2}, n-k}) \\ &= 1 - \frac{\alpha}{2} - \frac{\alpha}{2} \\ &= 1 - \alpha. \end{aligned}$$

Confidence Intervals

Proof 2: Confidence Interval Estimator Part 2

We know that $\frac{\hat{\beta}_j - \beta_j}{\text{se}[\hat{\beta}_j]}$ is a t-distributed random variable with $n - k$ degrees of freedom. So,

$$\begin{aligned} & \mathbb{P} \left(-t_{1-\frac{\alpha}{2}, n-k} \leq \frac{\hat{\beta}_j - \beta_j}{\text{se}[\hat{\beta}_j]} \leq t_{1-\frac{\alpha}{2}, n-k} \right) \\ &= \mathbb{P} \left(-\hat{\beta}_j - t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}[\hat{\beta}_j] \leq -\beta_j \leq -\hat{\beta}_j + t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}[\hat{\beta}_j] \right) \\ &= \mathbb{P} \left(\hat{\beta}_j - t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}[\hat{\beta}_j] \leq \beta_j \leq \hat{\beta}_j + t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}[\hat{\beta}_j] \right) \\ &= 1 - \alpha. \end{aligned}$$

Confidence Intervals

Proof 2: Confidence Interval Estimator Part 3

Thus, the probability that β_j lies in the interval $\hat{\beta}_j \pm t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}[\hat{\beta}_j]$ is $1 - \alpha$ meaning we can say with $100(1 - \alpha)\%$ confidence that β_j is some number in this interval. \square
Hooray!

Two-Sided Tests with Confidence Intervals

Property 2: Two-Sided Tests with Confidence Intervals

Suppose we wish to test $H_0 : \beta_j = a$ versus $H_A : \beta_j \neq a$ at the α significance level. If a is not contained in the confidence interval

$$\left[\hat{\beta}_j - t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}(\hat{\beta}_j), \hat{\beta}_j + t_{1-\frac{\alpha}{2}, n-k} \cdot \text{se}(\hat{\beta}_j) \right].$$

then we reject the null hypothesis.

F-Tests

Question 4

Suppose we are interested in the model

$$y_i = \mathbf{x}_i' \boldsymbol{\beta} + u_i.$$

How do we test $H_0 : \beta_{k-q} = 0 = \dots = \beta_k = 0$ versus $H_A : \beta_{k-q} \neq 0$ or \dots or $\beta_k \neq 0$?

Answer to Question 4

Construct an **F-Test**! How do we do that?

Unrestricted Model

Definition 9: Unrestricted Model

The **unrestricted model** is simply the full model:

$$y_i = \mathbf{x}'_i \boldsymbol{\beta} + u_i.$$

Restricted Model

Definition 9: Restricted Model

The **restricted model** assumes the last q parameters are zero so we get

$$y_i = \beta_0 + \beta_1 x_{i,1} + \dots + \beta_{k-q} x_{i,k-q} + u_i.$$

- The **restricted model** is “nested” within the **unrestricted model** because it has a subset of the covariates included in the unrestricted model.

Restricted and Unrestricted Model

Example 2

Suppose we have the **unrestricted model**

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} + u_i$$

and we wish to test $H_0 : \beta_1 = \beta_2 = 0$. Then, the **restricted model** is

$$y_i = \beta_0 + \beta_3 x_{i3} + v_i.$$

F-Statistic

Theorem 5: F-Statistic

The F-statistic

$$F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k)} \sim F_{q, n-k}$$

follows an F-distribution with q numerator degrees of freedom and $n - k$ denominator degrees of freedom.

- SSR_r is the SSR from the restricted model.
- SSR_{ur} is the SSR from the unrestricted model.
- q is the number of restrictions.
- $n - k$ is the number of estimated parameters (intercept included).
- Proof follows from the fact the F-statistic can be written as the ratio of two χ^2 random variables which necessarily follows an F-Distribution.

F-Test

Definition 10: F-Test

To conduct an **F-Test**, we:

1. Formulate a hypothesis test of $H_0 : \beta_{k-q} = \dots = \beta_k = 0$ versus $H_A : \beta_{k-q} \neq 0$ or \dots or $\beta_k \neq 0$.

F-Test

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2. Select a significance level α .

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2. Select a significance level α .
3. Estimate the unrestricted model to obtain the SSR_{ur} .

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2. Select a significance level α .
3. Estimate the unrestricted model to obtain the SSR_{ur} .
4. Estimate the restricted model to obtain the SSR_r .

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To conduct an **F-Test**, we:

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2. Select a significance level α .
3. Estimate the unrestricted model to obtain the SSR_{ur} .
4. Estimate the restricted model to obtain the SSR_r .
5. Construct the F-statistic $F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k)}$.

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2. Select a significance level α .
3. Estimate the unrestricted model to obtain the SSR_{ur} .
4. Estimate the restricted model to obtain the SSR_r .
5. Construct the F-statistic $F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k)}$.
6. Find the F-critical value $f_{1-\alpha, q, n-k}$ that corresponds to q numerator degrees of freedom and $n - k$ denominator degrees of freedom at the chosen significance level α .

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2. Select a significance level α .
3. Estimate the unrestricted model to obtain the SSR_{ur} .
4. Estimate the restricted model to obtain the SSR_r .
5. Construct the F-statistic $F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k)}$.
6. Find the F-critical value $f_{1-\alpha, q, n-k}$ that corresponds to q numerator degrees of freedom and $n - k$ denominator degrees of freedom at the chosen significance level α .
7. Reject the null hypothesis if and only if $F > f_{1-\alpha, q, n-k}$.

F-Test

Definition 11: F-Test Terminology

When we reject the null $H_0 : \beta_{k-q} = 0, \dots, \beta_k = 0$ we say these parameters are **jointly significant**. When we fail to reject this null, we say these parameters are **jointly insignificant**.

- When we reject this null hypothesis, we still do not know which of these variables effects the outcome significantly.
 - ▶ Perhaps they all do, but maybe only one does. So we must carry out **individual T-tests**.

F-Statistic for Intercept Only Model

Definition 11: F-Statistic for Intercept Only Model

The F-statistic for testing the null hypothesis $H_0 : \beta_1 = 0, \dots, \beta_k = 0$ is given by

$$F = \frac{R^2 / (k - 1)}{(1 - R^2) / (n - k)}.$$

- This is reported in R's `summary()` function.
- No need to estimate unrestricted and restricted models to carry out this test.

f-Critical Value

Definition 12: f-Critical Value

The **f-critical value** $f_{1-\alpha, q, n-k}$ for an F-test is the value such that the cumulative distribution function, $F_{f_{q, n-k}}$, of a F-distributed random variable, $f_{q, n-k}$, with q and $n - k$ degrees of freedom equals $1 - \alpha$:

$$\begin{aligned} F_{f_{q, n-k}}(f_{1-\alpha, q, n-k}) &= \mathbb{P}(f_{q, n-k} \leq f_{1-\alpha, q, n-k}) \\ &= 1 - \alpha. \end{aligned}$$

- We use the inverse CDF of the F-distribution to get $f_{1-\alpha, q, n-k}$.

p-value for F-Test

Example 3: p-value for F-Test

The p-value p corresponding to an F-test is given by

$$\begin{aligned} p &= \mathbb{P}(f_{q,n-k} > F) \\ &= 1 - F_{f_{q,n-k}}(F). \end{aligned}$$

where $f_{q,n-k}$ is a F-distributed random variable with degrees of freedom q and $n - k$.

- No need to multiply by two because the F-distribution only has a right tail.

Thank You!